

# Revisiting the Causal link Between Agriculture, Industrial output and Financial Sector Development in South Africa

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**Abstract:** This study seeks to establish the relationship between agriculture, industrial output and financial sector development in South Africa. It uses the Autoregressive Distributed Lag, Error Correction models and Granger causality techniques to test for long- and short-run relationships. The evidence from the models indicates the presence of a long-run relationship between industrial output and agriculture, which suggests that these sectors depend on each other for raw materials and inputs. In addition, stock market development represented by market capitalization has a long-run relationship with agriculture. However, no long-run relationship is established between credit extension and agriculture; and between gross fixed capital formation and agriculture, suggesting that an increase in agricultural output does not impact investment in long-term fixed assets. The evidence also shows a long-run relationship between exports and agricultural output, which is consistent with the export-led growth hypothesis. These findings have implications for policy formulation and allocation of resources.

**Keywords:** Granger Causality, agricultural output, industrial output, financial development.

## INTRODUCTION

This paper assesses the relationship between agriculture, industrialization and financial sector development in South Africa. South Africa has the most sophisticated and developed financial sector on the African continent. It is thus imperative to establish the nexus between agriculture, industrialization and financial sector development to inform and influence policy. Most economies are arguably highly dependent on agriculture for employment, including informal self-employment (Bruinsma, 2017; Davis, Di Giuseppe and Zezza, 2017). In addition, agriculture is an integral part of most countries' economies; inclusive growth should thus encapsulate the rural population's contributions to agriculture (Ding and Qian, 2016; Davis, Di Giuseppe and Zezza, 2017). As discussed below, it is argued both theoretically and empirically that agriculture influences economic development by positively affecting other non-agricultural sectors. We acknowledge that financial sector development has two legs – intermediary development and stock market development – and both are considered for completeness and comparison of the variables. Intermediary development captures the role played by

financial institutions in lending to the private sector, while stock market development captures the financial depth of the economy (Komal and Abbas, 2015).

Agriculture contributes approximately three per cent of South Africa's Gross Domestic Product (GDP). This mainly emanates from cattle and sheep farming as only 13 per cent of the land is suitable for crop farming. However, agro-processing makes a significant contribution to the amount generated by the manufacturing sector (Stats SA, 2018). Industrial output, represented by manufacturing, contributes approximately 13 per cent of GDP. Manufacturing is largely dominated by industries like automotive, chemicals, information and communication technology, electronics, metals, textiles and clothing and footwear. The financial sector's contribution (measured as finance, real estate and business services) to GDP is approximately 22 per cent. This is the largest sector in South Africa. In this study, we use stock market capitalization and domestic credit to the private sector as a proxy for financial sector development. The descriptive statistics are provided in the findings section.

## LITERATURE REVIEW

Discourse on financial development and economic development tracks back to Schumpeter (1911) who asserts that financial development is instrumental in

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JEL: G10; N57; O14; O16; Q11; Q14.

economic growth. Later in 1956, Solow identifies the main drivers of economic growth as the initial levels of gross domestic product, human capital and physical capital accumulation. The relationship between financial development and industrialization has been studied before (see Goldsmith 1969; McKinnon 1973; King and Levine, 1993; Rajan and Zingales 1998; Ayyagari, Demirgüç-Kunt, and Maksimovic, 2006). However, it remains imperative to revisit the nexus between these two sectors in the South African context given the rapid changes in both sectors.

Extant literature examines how agricultural productivity together with its associated exports led to increased economic growth through the income distribution model – a widely used economic model in macroeconomics (Galor and Zeira, 1993; Sihna, Pearson and Kadekodi, 2017). As argued by Murphy, Shleifer and Vishny (1989), theoretically, agricultural productivity is conducive for industrialization and economic development as (1) it is a source of autonomous demand for manufactured goods; and (2) income from agriculture is distributed to economic agents, resulting in increased demand for mass-produced domestic goods. Empirically, agricultural productivity led to increased domestic demand and industrialization in Japan (Rosenstein-Rodan, 1943) and boosted the initial stages of industrialization in postwar Taiwan (Ranis, 1979).

However, in Japan, banks concentrated on financing agriculture, domestic commerce and international trade during this period (Patrick, 1966). Thus, according to Patrick, the demand for financial services depends on the growth of real output and monetization of agriculture; hence, the demand following and supply leading phenomena. A recent study by Obilor (2013) asserts that credit to agriculture improves agricultural productivity. However, Obilor fails to articulate how this impacts on industrialization, banking sector development and economic development. In contrast, Cucinelli (2015) argues that crises in the agricultural sector can increase the level of non-performing loans in the economy as witnessed in the United States in the late 1970s.

On the other hand, Eddine Chebbi (2010) establishes that agricultural output influences other non-agricultural sectors in the short run, although agriculture itself does not benefit from the development of commerce and the services sectors. A key finding by Eddine Chebbi is that constraints on bank credit hamper growth in agricultural output. In a regression-

based study, Nnamocha and Eke (2015) confirm that bank credit and industrial output influence agricultural output in the long run. In addition, industrial output has short-run linkages with agricultural output. Bustos, Garber and Ponticelli (2016) analyzed the influence of agricultural productivity (with a focus on Genetically Engineered (GE) Soy seed) on bank deposits and credit and established that regions that engaged in GE Soy production experienced large bank deposits which did not, however, translate to increased credit by these branches. Interestingly, Bustos, Garber and Ponticelli found that bank branch networks led to increases in bank lending and larger firm growth in small to medium-size firms, which is in line with the theoretical underpinnings of Murphy, Shleifer and Vishny (1989) discussed earlier. According to Atteridge (2011), agriculture, particularly small-scale production that sustains local livelihoods, has been unattractive for private investment due to the inherent risk associated with such segments of the population. As a result, private investment has focused on large-scale export agricultural activities.

It should be noted, however, that the structure and composition of farming activities influence the impact of agricultural productivity. For instance, according to Murphy, Shleifer and Vishny (1989), in Columbia, a boom in tobacco exports in the 1850s and '60s did not lead to widespread economic development while the boom in coffee exports from 1880-1915 led to economic development. The distinguishing factor was that a few farmers planted tobacco on large tracts of land while coffee was planted on small tracts of land owned by many farmers. This subsequently led to income being widely spread among economic agents. With tobacco, a few farmers realized large amounts of income that they subsequently spent on luxury goods; hence, no economic impact was registered.

Several opportunities that arise from agricultural output have been identified by practitioners and scholars. For instance, it could result in enhanced communication and information systems (weather forecasting), off-the-grid energy, housing and construction, water management and financial services such as insurance (weather risk insurance, crop insurance and catastrophic bonds) (Pauw, 2015). Pauw concurs with the theory that agricultural output influences other economic sectors. Abbas and Choudhury (2013) further assert that agricultural output has direct causality with energy consumption, thus affirming agriculture's importance for other sectors of the economy.

The literature on the relationship between financial development and industrial output shows that industrial output is affected by levels of interest rates and bank credit (Olokoyo, Adetiloye and Ikpefan, 2016). Well-developed financial systems play an important role in innovation and improvement in the profitability of investments leading to improved industrial output (Shahbaz and Lean, 2012). Ductor and Grechyna (2015:5) argue, "Financial development can decrease economic growth if it is not accompanied by development in the real sector of the economy." Consequently, financial development should influence industrial output in order to contribute to economic growth. Ductor and Grechyna call this an interdependence relationship between financial sector development and industrialization.

Evidence shows that financial sector development influences industrial output (Wang, 2000). Using stock market development as a proxy for financial development, Udegbonam (2002) shows strong evidence suggesting that financial sector development influences industrial output. Fang and Miller (2014) concur and provide overwhelming evidence that volatility in financial development leads to industrial output volatility. Gokmenoglu, Ozatac and Eren (2015) show evidence of a long run equilibrium relationship between industrialization and financial development in Turkey although Granger causality tests showed no directional causality. Neusser and Kugler (1998) on the other hand, provide evidence of no co-integration between manufacturing GDP and financial development in OECD countries. There is no consensus on how financial development, agriculture and industrialization relate to each other. The next section discusses the source of data and methodology used in this study.

## DATA AND METHODOLOGY

The study uses secondary data from the McGregor BFA Library, Statistics South Africa (StatsSA) and the South African Reserve Bank (SARB) spanning 1994 (the year of the democratic transition) to September 2018. There are 99 quarterly observations for each series. A time series approach is employed for data analysis to determine whether long-run relationships exist between agricultural output and various sectors of the South African economy. The variables used are outlined below.

The main estimation variable is the natural logarithm of agricultural output (*Agric*). This is the

primary variable under investigation and it is assumed to contain all information and output from the agricultural sector. The second variable is a proxy for industrialization, which is measured by manufacturing output (*Ind*). The classification used by StatsSA is employed in the selection of this variable and it excludes construction. Financial development has two legs, intermediary development and stock market development, both of which are considered in this study. Intermediary development is represented by domestic credit extension to the private sector (*Cred*) while stock market development is represented by the JSE market capitalization (*Mkt*). Additional variables are included for the completeness of the estimation model below. These are gross fixed capital formation (GCF), exports (*Exp*), and the non-agriculture unemployment rate (UER).

The Autoregressive Distributed Lag (ARDL) cointegration technique recommended by Pesaran and Pesaran (1997), Pesaran and Shin (1998) and Pesaran, Shin and Smith (2001) is used to test for a long-run relationship between agricultural output, industrial output and financial development in South Africa. The advantage of using the ARDL estimation technique is that, unlike other techniques such as the vector error correction model, it does not require that the variables be integrated of the same order or that they be stationary or non-stationary (Hamuda, Šulíková, Gazda and Horváth, 2013; Asumadu-Sarkodie and Owusu, 2016). Therefore, the following basic model was estimated:

$$\Delta Y_t = \beta_0 + \sum_{i=1}^n \beta_i \Delta y_{t-i} + \sum_{i=0}^n \delta_i \Delta x_{t-i} + \varphi_1 y_{t-1} + \varphi_2 x_{t-1} + \mu_t \quad (i)$$

Where:  $\beta_i$  and  $\delta_i$  are the short-run coefficients,  $\varphi_1$  and  $\varphi_2$  are the ARDL long-run coefficients and  $\mu_t$  is the white noise term. The model specification based on the estimation model above is as follows:

$$\Delta \ln(Ind)_t = \beta_0 + \sum_{i=1}^n \beta_i \Delta \ln(Ind)_{t-i} + \sum_{i=0}^n \delta_i \Delta \ln(Agric)_{t-i} + \varphi_1 \ln(Ind)_{t-1} + \varphi_2 \ln(Agric)_{t-1} + \mu_t \quad (ii)$$

Where:  $\beta_i$  are the short-run coefficients with  $\Delta$  as the first-difference operator. The  $\varphi_i$  are the long-run coefficients, which were used to test for long-run causality between *Agric* and other variables in the series. *Ind* is replaced by the other variables for subsequent estimations. The estimation test hypothesis is that:

$$H_0 : \varphi_1 = \varphi_2 = 0 \text{ against } H_1 : \varphi_1 = \varphi_2 \neq 0.$$

The bounds tests can either show the presence of a long-run relationship or the absence thereof. In the case of the former, an error correction model, as specified in Equation *iii* below is estimated to capture the short-run dynamics of the long-run relationship such as the speed of adjustment back to equilibrium. In the case of the latter, a short-run model as specified in Equation *iv* is estimated.

$$\Delta \ln(Ind)_t = \vartheta_0 + \vartheta_1 \Delta \ln(Ind)_{t-1} + \vartheta_2 \Delta \ln(Agric)_{t-1} + \pi ECT_{t-1} + \varepsilon_t \tag{iii}$$

$$\Delta \ln(Ind)_t = \theta_0 + \theta_1 \Delta \ln(Ind)_{t-1} + \theta_2 \Delta \ln(Agric)_{t-1} + \varepsilon_t \tag{iv}$$

Where: The lagged first differenced values on the explanatory variables show the short-run causality between *Agric* and *Ind*. To denote the long-run causality, the estimation must yield a negative and statistically significant coefficient on the ECT. The same approach is applied for *Agric* and the other variables. Pairwise Granger-causality tests were also estimated to augment the short-run analysis.

## FINDINGS

### Descriptive Statistics

Table 1 presents the results of the descriptive statistics. Average manufacturing output, domestic credit to the private sector, JSE market capitalization, gross fixed capital formation, exports and non-agricultural unemployment rate were R293 493 000, R4 453 906 000, R5 478 466 000, R439 489 000, R736 324 000 and 24.08 %, respectively, when average agricultural output was R53 692 million for the period 1994 to 2018. When maximum agricultural output was R115 550 000, maximum manufacturing output, domestic credit to private sector, JSE market

capitalization, gross fixed capital formation, exports and non-agricultural unemployment rate were R604 623 000, R10 740 153 000, R15 294 514 000, R652 624 000, R974 672 000 and 29.3%, respectively.

It is interesting to note that the non-agricultural unemployment rate increased as agricultural output increased. This is possible because the agricultural industry is becoming increasingly capital intensive and high output levels could accrue due to efficiency arising from mechanization. Table 1 also shows that most of the variables are fairly symmetrical given the skewness values of between -0.5 and 0.5, except for the non-agricultural unemployment rate and JSE market capitalization which are moderately skewed. The non-agricultural unemployment rate is negatively skewed, suggesting very few low values on average and JSE market capitalization is positively skewed, suggesting very few high values. Overall, all the variables show that there were a few outliers from the data set. This finding is supported by the negative kurtosis measure, which shows that the data sets were light-tailed.

### Stationarity and Unit Root Tests

Stationarity and unit root tests were also conducted and are shown in Table 2. All the variables in this study were either I(0) or I(1), confirming the applicability of the ARDL model which requires that all variables are below I(2) (Arshed, 2014). Therefore, the ARDL model was used to test the long-run relationship between agricultural output and manufacturing output, domestic credit to private sector, JSE market capitalization, gross fixed capital formation, exports and the non-agricultural unemployment rate.

### ARDL Model Estimation

Prior to estimating the ARDL models for *Agric* and the six variables, the number of optimum lags and the

Table 1: Descriptive Statistics

Statistic	Agric	Cred	Exp	Ind	GCF	UER	Mkt
Mean	53692	4453906	736324	293493	439489	24,08	5478466
Median	44708	3746800	752099	271037	451850	24,50	4464192
Std. Dev	26485	3242203	141425	143070	150887	2,77	4679833
Kurtosis	-0,87	-1,24	-0,97	-0,93	-1,71	0,31	-0,66
Skewness	0,49	0,40	-0,31	0,38	-0,06	-0,69	0,83
Minimum	15983	500333	413891	92269	209074	16,90	786248
Maximum	115550	1074015	974672	604623	652624	29,30	15294514

**Table 2: Stationarity and Unit Root Tests**

ADF	I(0)		I(1)		Order
	C	C + @trend	C	C + @trend	
LNAGRIC	0.8357	0.0006***	0.0001***	0.0000***	I(0)
LNIND	0.5043	0.4352	0.0000***	0.0000***	I(1)
LNMKT	0.0835*	0.9560	0.0000***	0.0000***	I(1)
LNCRED	0.6253	0.0857*	0.0001***	0.0000***	I(1)
LNGCF	0.4639	0.9222	0.0000***	0.0000***	I(1)
LNEXP	0.4572	0.0239**	0.0001***	0.0000***	I(0)
UER	0.1721	0.2231	0.0000***	0.0000***	I(1)
KPSS	I(0)		I(1)		Order
	C	C + @trend	C	C + @trend	
LNAGRIC	1.3206***	0.0732	0.1602	0.1601**	I(1)
LNIND	1.3279***	0.2826***	0.2569	0.0593	I(1)
LNMKT	1.2983***	0.3106***	0.5979**	0.0418	I(1)
LNCRED	1.3179***	0.2277***	0.1564	0.1203*	I(1)
LNGCF	1.2811***	0.1771**	0.2999	0.1137	I(1)
LNEXP	1.2511***	0.2358***	0.2394	0.2000**	I(1)
UER	0.5182**	0.1395*	0.0909	0.0877	I(1)

best ARDL model to estimate were determined using the SBIC, AIC and HQIC. The same procedure is also followed in the determination of the relationship between Ind and the two financial development measures – Cred and Mkt. The maximum number of lags selected was four for all six models. As shown in Table 3, ARDL(2,0) was selected for Agric vs Ind, Cred, GCF and Exp. For Agric vs Mkt and UER, ARDL(4,3) and ARDL(1,0) were selected, respectively. ARDL(2,0) was selected for both Ind vs Cred and Ind vs Mkt. Subsequent to the confirmation of the selected models, diagnostic tests were conducted on these models. The results from the tests showed that there was no serial autocorrelation and heteroscedasticity in

any of the six models. Model stability tests – CUSUM and CUSUM of squares tests – also showed that the models were stable. Therefore, the bounds tests for cointegration were subsequently conducted.

As shown in Table 4 below, based on the lower and upper bound critical values provided by Pesaran *et al.* (2001), the null hypothesis that there is no cointegration was rejected at 5% level of significance for Ind vs Agric, Mkt vs Agric and Exp vs Agric. All the estimated F-values (11.39496, 6.017596 and 5.943577, respectively) were greater than the upper bound critical value (4.16). The same applies for Ind vs Cred (10.62484) and Mkt (11.70374). This suggests

**Table 3: Lag length and ARDL Model Selection**

	Variables	Max lags	ARDL model
Model 1	Ind vs Agric	4	ARDL(2,0)
Model 2	Cred vs Agric	4	ARDL(2,0)
Model 3	Mkt vs Agric	4	ARDL(4,3)
Model 4	GCF vs Agric	4	ARDL(2,0)
Model 5	Exp vs Agric	4	ARDL(2,0)
Model 6	UER vs Agric	4	ARDL(1,0)
Model 7	Ind vs Cred	4	ARDL(2,0)
Model 8	Ind vs Mkt	4	ARDL(2,0)

that there is a long-run relationship between Agric and these three variables as well as between Ind and the financial development measures, results that warrant the estimation of error correction models. The presence of a long-run relationship between Ind and Agric suggests that these sectors depend on each other for raw materials and inputs. Furthermore, as suggested by Murphy, Shleifer and Vishny (1989), agricultural productivity is a source of autonomous demand for manufactured goods, and income from agriculture leads to increased demand for mass-produced domestic goods. Manufacturing output thus increases to meet this demand.

The pattern seems to match those in Japan (Rosenstein-Rodan, 1943) and postwar Taiwan (Ranis, 1979). The long-run relationship is also significant for the Mkt and Agric. This is in line with Bustos, Garber and Ponticelli's finding that bank branch networks lead to increases in bank lending and larger firm growth in small- to medium-size firms that, in turn, increases market development. Furthermore, an increase in agricultural output can have an impact on agricultural derivatives trading, a component of financial markets that has grown significantly over the years. For Exp and Agric, the long-run relationship is consistent with the export-led growth hypothesis proposed by Galor and Zeira (1993) and Sihna, Pearson and Kadekodi (2017). Demand for exports may lead to the need for increased production and output in various sectors, including agriculture.

For Ind and Cred and Ind and Mkt, the significant relationship shows that financial market development contributes significantly to industrial growth and output. Per Fisman and Love (2003), financial market development enables the channelling of funds through bank lending, market-based lending and equity

fundraising to support industrial development. These funds can be domestic funds from domestic savings or international funds through the foreign portfolio and foreign direct investment. The investment in plant and equipment, research and development and the expansion of businesses increase industrial output. Beck and Levine (2002) also reported that overall financial development boost industry growth, new establishment formation, and efficient capital allocation. Further, investment and development of other supportive infrastructure for industrial development usually come about as a result of proper financial market development (Demirgüç-Kunt and Levine, 2004).

For Cred vs Agric, GCF vs Agric and UER vs Agric, the null hypothesis of no long-run relationship could not be rejected at the 5% level of significance. This is because the estimated F-values (3.524962, 2.783247 and 2.849273, respectively) were all lower than the lower bound critical value (3.62). As such, there was no need to estimate error correction models for these variables. Instead, short-run models were estimated to examine the short-run relationship between Agric and these variables. The absence of a long-run relationship between Cred and Agric matches Bustos, Garber and Ponticelli's (2016) findings. These researchers analyzed the influence of agricultural productivity on bank deposits and credit and established that regions engaged in GE Soy production experienced large bank deposits, which did not, however, translate to increased credit by these branches.

For GCF and Agric, the absence of a long-run relationship seems to suggest that an increase in agricultural output does not impact investment in long-term fixed assets. Conceivably, once-off purchases of fixed assets in the agricultural sector can increase for a

**Table 4: Bounds Test for Cointegration**

	Estimated F-value	Pesaran critical values at 5%	
		Lower bound	Upper bound
Ind vs Agric	11.39496	3.62	4.16
Cred vs Agric	3.524962	3.62	4.16
Mkt vs Agric	6.017596	3.62	4.16
GCF vs Agric	2.783247	3.62	4.16
Exp vs Agric	5.943577	3.62	4.16
UER vs Agric	2.849273	3.62	4.16
Ind vs Cred	10.62484	3.62	4.16
Ind vs Mkt	11.70374	3.62	4.16

**Table 5: Error Correction Models**

	Error term	t-statistic	Prob.*
Ind vs Agric	-0.502727	-1.709457	0.0907
Mkt vs Agric	-0.081187	-2.518478	0.0136
Exp vs Agric	-0.903287	-3.429484	0.0009
Ind vs Cred	-0.130404	-5.706138	0.0000
Ind vs Mkt	-0.285159	-5.988847	0.0000

number of years without any need for new purchases. Furthermore, the increased output may produce substantial funds, but these may be directed towards repayment of loans rather than investment in new fixed assets. For UER, it is likely that increased agricultural output comes from mechanization and capital-intensive developments such as genetic modification of seeds that do not impact on employment levels in South Africa. In addition, the increased agricultural output does not seem to be large enough to impact employment in other, non-agricultural sectors.

### Error Correction Models

Error correction models were estimated for the models that yielded significant cointegrating relationships between Agric and the respective variables. Table 5 reports the results from these estimations. Similar diagnostic tests to those conducted in the ARDL estimations were again conducted before any inferences could be made from the results. All three models passed these tests. The error term coefficient was found to be negative and statistically significant in all three cases, although it was significant at 10% for the Ind vs Agric model. The results show that about 50%, 8% and 90% of any disequilibrium between Ind and Agric, Mkt and Agric and Exp and Agric, respectively, is corrected in each quarter. This means that a change in agricultural output would take 1.99 quarters ( $1/0.502727$ ), 12.32 quarters ( $1/0.081187$ ) and 1.11 quarters ( $1/0.903287$ ), respectively, for a correction back to equilibrium to occur. Thus, Exp and Ind are quick to correct back to equilibrium while Mkt is rather slow in its correction. It is conceivable given that industrial output and exports are economically more linked to agricultural output than stock market development. For Ind and Cred and Ind and Mkt, the correction to equilibrium is faster for the latter relative to the former, but the coefficients are both negative and statistically significant. It takes 7.67 quarters ( $1/0.130404$ ) for Ind and Cred while it takes 3.51 quarters ( $1/0.285159$ ) for Ind and Mkt to get back

to equilibrium. This seems to suggest that the primary source of financing that leads to a greater form of expansion and increased industrial output is equity relative to debt, albeit both significant.

### Short-Run Models

Short-run models were estimated for the variables that yielded no significant long-run cointegrating relationships with Agric, namely, Cred, GCF and UER. The results are reported in Table 6. The diagnostic tests were also conducted on the three models and they passed all the tests. The results show that there is no significant short-run relationship between Agric and all three variables as seen by the insignificant coefficients on Agric in all three estimations. One-period lagged changes in Cred and GCF influence the current values of the same variables, respectively. However, for UER, the effect is insignificant.

### Granger-Causality Tests

The short-run relationship between Agric and the six variables was further examined using the Granger-causality pairwise tests. Table 7 below exhibits the results from the estimations. There is evidence that Ind, Cred, Mkt and Exp Granger-cause Agric as all the null hypotheses that these variables do not Granger-cause Agric can be rejected in favour of the alternative. However, the causality is unidirectional as the null hypothesis that Agric does not Granger-cause any of these variables cannot be rejected. This finding seems to suggest that an increase in manufacturing output, the extension of credit to the private sector, stock market development and exports all cause an increase in agricultural output.

For GCF and UER, there is no Granger-causality with Agric as none of the null hypotheses can be rejected. This suggests that an increase in agricultural output does not impact gross fixed capital formation and unemployment and vice versa. The absence of causality between agriculture and unemployment could

**Table 6: Short-Run Models**

	Coefficient	Std. Error	t-Statistic	Prob.
Cred vs Agric				
C	0.046170	0.021330	2.164517	0.0330
D(LNCRED(-1))	0.635973	0.077883	-8.165754	0.0000
D(LNAGRIC(-1))	0.201337	0.225252	0.893831	0.3737
GCF vs Agric				
C	0.004497	0.002211	2.034024	0.0448
D(LNGCF(-1))	0.515565	0.087987	5.859569	0.0000
D(LNAGRIC(-1))	0.033637	0.020876	1.611300	0.1105
UER vs Agric				
C	0.088325	0.127096	0.694949	0.4888
D(UER(-1))	0.119440	0.102612	-1.164000	0.2474
D(LNAGRIC(-1))	0.133491	1.341927	0.099477	0.9210

**Table 7: Pairwise Granger-Causality Tests**

Null Hypothesis:	Obs	Prob.	Causality
LNIND does not Granger Cause LNAGRIC	95	0.0050	Yes
LNAGRIC does not Granger Cause LNIND		0.8747	No
LNCRED does not Granger Cause LNAGRIC	95	0.0030	Yes
LNAGRIC does not Granger Cause LNCRED		0.6312	No
LNMTKT does not Granger Cause LNAGRIC	95	0.0143	Yes
LNAGRIC does not Granger Cause LNMTKT		0.1362	No
LNGCF does not Granger Cause LNAGRIC	95	0.3963	No
LNAGRIC does not Granger Cause LNGCF		0.4725	No
LNEXP does not Granger Cause LNAGRIC	95	0.0160	Yes
LNAGRIC does not Granger Cause LNEXP		0.9779	No
UER does not Granger Cause LNAGRIC	95	0.8870	No
LNAGRIC does not Granger Cause UER		0.3791	No
LNMTKT does not Granger Cause LNIND	95	0.3769	No
LNIND does not Granger Cause LNMTKT		0.0819	Yes
LNCRED does not Granger Cause LNIND	95	0.6899	No
LNIND does not Granger Cause LNCRED		0.1169	No

be due to increased capital-intensive automation in agriculture that reduces the need for labour as well as movement to jobs other than agriculture as the country becomes increasingly affluent. The lack of causality between agriculture and gross fixed capital formation could be because much investment is being channelled to the secondary and tertiary sectors rather than the primary sector under which agriculture falls.

For Ind and the two financial market development measures, the only significant causality (at 10%) runs from Ind to Mkt. There is no evidence of causality

between Ind and Cred. This is to be expected as financial market development may take longer to impact industrial output and growth. This is why there is a significant long run relationship between Ind and the two financial development measures while the short run causality is either weak or absent.

## CONCLUSION

This study sought to establish the long run cointegrating relationships and short run causal linkages between agriculture, industrial output and

financial sector development in South Africa. The evidence from the models indicates the presence of a long-run relationship between industrial output and agriculture, which suggests that these sectors depend on each other for raw materials and inputs. In addition, stock market development represented by market capitalization has a long-run relationship with agriculture. However, there is no long-run relationship between credit extension and agriculture; and between gross fixed capital formation and agriculture, suggesting that an increase in agricultural output does not impact investment in long-term fixed assets. Industrial output and growth have a significant long run relationship with financial market development measures but poor or absent short run causality linkage. This is because financial market development takes a long time to impact industrial output and growth. Future studies should consider disaggregating industrial output to determine the linkages of specific sectors with agriculture and financial sector development.

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Received on 01-10-2019

Accepted on 13-11-2019

Published on 27-12-2019

DOI: <https://doi.org/10.6000/1929-7092.2019.08.109>© 2019 Sibanda *et al.*; Licensee Lifescience Global.

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